# Mathematical Statistics

## SinkingFun

## November 2018

### Hi!

I'm currently preparing for a final test on Mathematical Statistics and as you might have guessed, my doubts are more on the theoretical part of the course.

So, I have worked on three different questions but have not been able to finish any of them, here is what I 've done so far:

## Problem 1

1. The sum of the values obtained in a random sample of size 5 taken from a population with Poisson distribution will be used to test the null hypothesis

 $H_0$ : population mean > 2

against the alternative hypothesis:

 $H_1$ : population mean  $\leq 2$ 

The null hypothesis will be rejected if and only if the sum of the observations is 5 or less.

• Determine:

The power function for the means: 1.5, 2, 2.4, 2.6 and

Obtain a table with probabilities of errors type 1 and 2

#### My solution:

Let  $X \sim \text{Poisson}(\theta)$ , so that  $E[X] = \theta$ 

Then let  $X_i$  with i in  $\{1, 2, \ldots, 5\}$  be the random sample, so

 $Y = \sum_{i=1}^{5} X_i$ , and we reject  $H_0$  iff  $Y \leq 5$ 

Then, we know that  $Y \sim \text{Poisson}(5\theta)$  and so our power function will be given by:

$$\begin{aligned} \pi(\theta) &= \mathbf{P}(Y \le 5) = \sum_{y=0}^{5} e^{-5\theta} \frac{(5\theta)^y}{y!} \\ \text{i.e. } \pi(\theta) &= e^{-5\theta} (1 + 5\theta + \frac{(5\theta)^2}{2!} + \frac{(5\theta)^3}{3!} + \frac{(5\theta)^4}{4!} + \frac{(5\theta)^5}{5!}) \end{aligned}$$

And from here, the first part of the problem would be to get the values of

 $\pi(1.5), \pi(2), \pi(2.4)$  and  $\pi(2.6)$  right?

But what I don't get in this problem is how to construct the *classical* table with probabilities of errors Type I and II.

## Then, for Problem 2

2. Consider a population with Bernoulli distribution, we have the hypothesis test:

 $H_0: \mathbf{p} = p_0$  $H_1: \mathbf{p} < p_0$  $\alpha$ 

Given a sample of size "n", use the correct theory to find the region of rejection, the test statistic (for the exact test and for a large sample size), and the property that the region has in order to generate the corresponding unilateral alternative hypothesis and specify it

#### My solution

I followed the process of a similar problem that we saw on class, so here we start by considering  $X \sim Bernoulli(p)$  and we note that  $\hat{p} = \frac{X}{n}$ 

Then, our likelyhood function is:

• 
$$L(x_1, ..., x_n) = {n \choose X} p^X (1-p)^{n-X}$$
 with  $X \in 0, ..., n$ 

we let  $\omega_0 = [p_0, 1)$  and  $\omega_1 = (0, p_0)$  so that:

•  $\Omega = \omega_0 \cup \omega_1 = (0, p_0) \cup [p_0, 1) = (0, 1)$ 

And using the likelyhood ratio we get:

• 
$$\Lambda = \frac{ML_0}{ML} = \frac{p_0^X (1-p_0)^{n-X}}{p^X (1-p)^{n-X}} \le k_1$$
 with  $k_1 \in (0,1)$ 

Taking  $ln(\Lambda) = h(\hat{p})$  we get:

• 
$$h(\hat{p}) = \hat{p}ln(p_0) + (1-\hat{p})ln(1-p_0) - \hat{p}ln(\hat{p}) - (1-\hat{p})ln(1-\hat{p}) \le k$$
 with  $k < 0$ 

We then analyze the behavior of  $h(\hat{p})$  with the following 2 limits:

- $\lim_{\hat{p}\to 0} h(\hat{p}) = ln(1-p_0)$ , and
- $\lim_{\hat{p} \to 1} h(\hat{p}) = \ln(p_0)$

And taking the derivative we get:

• 
$$h(\hat{p})' = ln(\frac{p_0(1-\hat{p})}{\hat{p}(1-p_0)}) \ge 0 \iff \frac{p_0(1-\hat{p})}{\hat{p}(1-p_0)} \ge 1 \iff p_0(1-\hat{p}) \ge \hat{p}(1-p_0)$$

hence

• 
$$p_0 \ge \hat{p}$$

But from here I don't know how to continue. I think the next step is to give a size of an arbitrary  $\alpha$ , i.e. some sort of  $P(X \leq somebound | p = p_0) = \alpha$  in order to find the unilateral interval, but I don't know how to get there.

Also, I think the wanted test statistic is  $Z = \frac{\hat{p} - p_0}{\sqrt{\frac{p \cdot p_0}{n}}}$  but I'm also not sure.

## And finally, Problem 3

3. Given a random sample of size "n" taken from a normally distributed population with unknown mean and variance, find the region of rejection and the test statistic (completely specified) to prove the null hypothesis

 $H_0: \sigma = \sigma_0$  against the alternative hypothesis:

 $H_1: \sigma \neq \sigma_0$ 

#### My Solution

I'm actually sort of lost in this one, the only thing I think I've figured out is that if we consider  $X_i \sim N(\mu, \sigma)$  independent, we then can get

•  $\frac{nS_n^2}{\sigma^2} = \sum_{i=1}^n \left(\frac{(x_i - \bar{x})^2}{\sigma^2}\right)$ 

and then, by the m.g.f, we know that:

•  $\sum_{i=1}^{n} \left( \frac{(x_i - \bar{x})^2}{\sigma^2} \right) \sim \chi_{n-1}^2$  so that  $\frac{nS_n^2}{\sigma^2} = \chi_{n-1}^2$ 

And from there I don't know how to continue.

So, I know it is sort of a long post, and I thank you for the time you take on reading this. Any help you can provide will be great for a starting-to-panic student.